

Blind Signal Separation & Multivariate Non-Linear Dependency Measures

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Outline: a package of (2) methodologies ☺

- 1 Part I: Causal dependencies in multivariate time series
 - Time series graphs & information theoretic measure
- 2 Part II: Tracking dependency in large multivariate financial systems.
 - Coupling & decoupling information



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SYstemic Risk **TO**mography

*Signals, Measurements, Transmission Channels,
and Policy Interventions*

Part I: Multivariate Dependency Measures

“The kiss of information theory that captures systemic risk tomography” (joint with Philippe de Peretti)

- Detection and quantification of **causal dependencies** in multivariate time series
- How do we create a formal measure of systemic risk that adequately **captures complex linkages** in the financial system? 😊
- Granger causality & transfer entropy ??



Drawbacks of Granger causality & transfer entropy ☹️

- Transfer entropy is the **information-theoretic analogue** of Granger causality
- It reduces to Granger causality for vector auto-regressive processes
- It is advantageous **for the analysis of non-linear signals** where the model assumption of Granger causality might not hold.
- The transfer entropy is **not uniquely determined** by the interaction of the two components alone and **depends on misleading effects** such as autodependency and interaction with other process.
- It requires arbitrary truncation during **estimation**, it usually requires more samples for accurate estimation
- Can lead to false interpretation since it is **not lag-specific**.

Our Offer 😊

- A novel **information theoretic** model-free approach to understanding systemic risk
- To detect and quantify **causal dependencies** from multivariate time series.
- It gives **similar scores to equally noisy dependencies**.
- It is uniquely determined by the **interaction of the two components alone** and in a way autonomous of their interaction with the remaining process.
- **Excludes the misleading influence** of autodependency within a process
- **Lag-specific**. Enhance better interpretation.
- Only requires that the multivariate time series be **stationary**.

Methodology

Let \mathfrak{X} be a multivariate time series with a set of subprocesses V at each time $t \in \mathbb{Z}$ and directional links be defined in E . Then

$$\mathcal{G} = (V \times \mathbb{Z}, E)$$

is the **time series graph** of \mathfrak{X} , where the set of nodes in the graph are made up of V .

- Like graphical models (Lauritzen 1996), TSG's are based on the **concept of conditional independence**.
- Note that the time-dependence in the time series is used to define directional links in the graph.

Definition (Lag-specific directed link)

We say that the nodes $\mathcal{A}_{t-\tau} \in \mathcal{G}$ and $\mathcal{B}_t \in \mathcal{G}$ are connected by a **lag-specific directed link** " $\mathcal{A}_t \longrightarrow^\tau \mathcal{B}_t$ " pointing forward in time if and only if $\tau > 0$ and

$$I_{\mathcal{A} \xrightarrow{\text{link}} \mathcal{B}}(\tau) \equiv I(\mathcal{A}_{t-\tau}; \mathcal{B}_t \mid \mathfrak{X}^- \setminus \{\mathcal{A}_{t-\tau}\}) > 0. \quad (1)$$

Thus, $\mathcal{A}_{t-\tau}$ and \mathcal{B}_t are **connected** if they are not independent conditionally on the past of the whole process **excluding** $\{\mathcal{A}_{t-\tau}\}$ (denoted by the symbol \setminus) which **implies a lag-specific causality** with respect to \mathfrak{X} . $I(\cdot; \cdot \mid \cdot)$ denotes conditional mutual information.

- If $\mathcal{A} \neq \mathcal{B}$ then " $\mathcal{A}_t \longrightarrow^\tau \mathcal{B}_t$ " represents a **coupling at lag** τ .
- An **autodependency** at lag τ corresponds to $\mathcal{A} = \mathcal{B}$.

Definition (Undirected contemporaneous link)

The nodes $\mathcal{A}_t \in \mathcal{G}$ and $\mathcal{B}_t \in \mathcal{G}$ are connected by an **undirected contemporaneous link** " $\mathcal{A} - \mathcal{B}$ " if and only if

$$I_{\mathcal{A}-\mathcal{B}}^{link} \equiv I(\mathcal{A}_t; \mathcal{B}_t \mid \mathfrak{X}_{t+1}^- \setminus \{\mathcal{A}_t, \mathcal{B}_t\}) > 0. \quad (2)$$

Definition (Notion of *parents* and *neighbors* of subprocesses.)

Given the nodes $\mathcal{A}_t \in \mathcal{G}$ and $\mathcal{B}_t \in \mathcal{G}$, the *parents* $\mathcal{P}_{\mathcal{B}_t}$ and *neighbors* $\mathcal{N}_{\mathcal{B}_t}$ of node \mathcal{B}_t are defined as

$$\mathcal{P}_{\mathcal{B}_t} \equiv \{\mathcal{A}_{t-\tau} : \mathcal{A} \in \mathfrak{X}, \tau > 0, \mathcal{A}_{t-\tau} \longrightarrow \mathcal{B}_t\} \quad (3)$$

$$\mathcal{N}_{\mathcal{B}_t} \equiv \{\mathcal{A}_t : \mathcal{A} \in \mathfrak{X}, \mathcal{A}_t - \mathcal{B}_t\} \quad (4)$$

Definition (Causal Markov Condition)

Any node $\mathcal{B}_t \in \mathcal{G}$ in the time series graph is **conditionally independent** of $\mathfrak{X}_t^- \setminus \mathcal{P}_{\mathcal{B}_t}$ given its *parents* $\mathcal{P}_{\mathcal{B}_t}$.

Definition (Momentary Information Transfer (MIT) Links)

For two subprocesses \mathcal{A} , \mathcal{B} of a stationary multivariate discrete time process \mathfrak{X} with parents $\mathcal{P}_{\mathcal{A}_t}$ and $\mathcal{P}_{\mathcal{B}_t}$ in the associated time series graph and $\tau > 0$, the general **information theoretic measure** between nodes $\mathcal{A}_{t-\tau}$ and \mathcal{B}_t is given by

$$\begin{aligned} I_{\mathcal{A} \xrightarrow{MIT} \mathcal{B}}(\tau) &\equiv I(\mathcal{A}_{t-\tau}; \mathcal{B}_t \mid \mathcal{P}_{\mathcal{B}_t} \setminus \{\mathcal{A}_{t-\tau}\}, \mathcal{P}_{\mathcal{A}_{t-\tau}}) > 0 \\ &= H(\mathcal{B}_t \mid \mathcal{P}_{\mathcal{B}_t} \setminus \{\mathcal{A}_{t-\tau}\}, \mathcal{P}_{\mathcal{A}_{t-\tau}}) - H(\mathcal{B}_t \mid \mathcal{P}_{\mathcal{B}_t}) \end{aligned}$$

and contemporaneous MIT defined by

$$I_{\mathcal{A} - \mathcal{B}}^{MIT} \equiv I(\mathcal{A}_t; \mathcal{B}_t \mid \mathcal{P}_{\mathcal{B}_t}, \mathcal{P}_{\mathcal{A}_t}, \mathcal{N}_{\mathcal{A}_t} \setminus \{\mathcal{B}_t\}, \mathcal{N}_{\mathcal{B}_t} \setminus \{\mathcal{A}_t\}, \mathcal{P}_{\mathcal{N}_{\mathcal{A}_t} \setminus \{\mathcal{B}_t\}}, \mathcal{P}_{\mathcal{N}_{\mathcal{B}_t} \setminus \{\mathcal{A}_t\}}). \quad (5)$$

where $H(\mathcal{X})$ is Shannon's entropy and $H(\mathcal{X} \mid \mathcal{Y})$ denotes conditional Shannon's entropy.

The *parents* of all subprocesses in \mathfrak{X} together with the contemporaneous links **forms the time series graph**.

Definition (Coupling Strength)

The partial correlation MIT measure, denoted ρ^{MIT} , associated with equation (5), for the **strength of coupling mechanism** between $\mathcal{A}_{t-\tau}$ and \mathcal{B}_t is given by

$$\rho_{\mathcal{A} \xrightarrow{MIT} \mathcal{B}}(\tau) \equiv \rho\left(\mathcal{A}_{t-\tau}; \mathcal{B}_t \mid \mathcal{P}_{\mathcal{B}_t} \setminus \{\mathcal{A}_{t-\tau}\}, \mathcal{P}_{\mathcal{A}_{t-\tau}}\right). \quad (6)$$

- The measure ρ^{MIT} quantifies **how much the variability** in \mathcal{A} at the exact lag τ directly influences \mathcal{B} , irrespective of the past of $\mathcal{A}_{t-\tau}$ and \mathcal{B}_t .
- ρ^{MIT} is the cross-correlation of the residual after $\mathcal{A}_{t-\tau}$ and \mathcal{B}_t have been regressed on both the parents of $\mathcal{A}_{t-\tau}$ and \mathcal{B}_t .
- The contemporaneous MIT in the linear case is equivalent to the partial correlation of the errors after regressing each process on its *parents*.
- Unlike classical statistics, interactions in the framework of information theory are **viewed as transfers of information** and thus this approach is model-free.

Estimation

Summary

- In the first Step, PC algorithm (Spirtes et al [SSCR 1991]) is used to **estimate the parents of each process**, i.e., as a variable selection method.
 - Unlike graphical models, only undirected links are inferred and the second step of PC algorithm is omitted.
 - This first step **determines the existence or absence of a link**, which also provide useful information on the causality between lagged components of the multivariate process.
- In the second step, **MIT is used and all possible links are tested** again.
 - In this step, the **problem of serial dependencies is drastically reduced** using MIT (Runge et al[Physical Review E, 2012]).

Simulation Exercise

Consider a simulated 1000 points of a **stationary** multivariate autoregressive process made up of four subprocesses $\{X_t, Y_t, Z_t, W_t\}'$ defined by

$$X_t = aX_{t-1} + cZ_{t-4} + \varepsilon_x \quad (7)$$

$$Y_t = kX_{t-1} + hY_{t-1} + \varepsilon_y \quad (8)$$

$$Z_t = dY_{t-2} + bZ_{t-1} + fW_{t-1} + \varepsilon_z \quad (9)$$

$$W_t = eY_{t-3} + gW_{t-1} + \varepsilon_w \quad (10)$$

and the innovation covariance matrix given by $\Sigma_\varepsilon = \begin{pmatrix} 1 & 0 & d & 0 \\ 0 & 1 & 0 & d \\ d & 0 & 1 & 0 \\ 0 & d & 0 & 1 \end{pmatrix}$, where

$a = 0.6, b = 0.4, c = 0.3, d = -0.3, e = -0.6, f = 0.2, g = 0.4, k = 0.3$, and $h = 0.6$.

Notice that the **lagged causal chain** for this process is $X \rightarrow^1 Y \rightarrow^2 Z$ with feedback $Z \rightarrow^4 X$, and $Y \rightarrow^3 W \rightarrow^1 Z$, plus contemporaneous links $X - Z$ and $Y - W$.

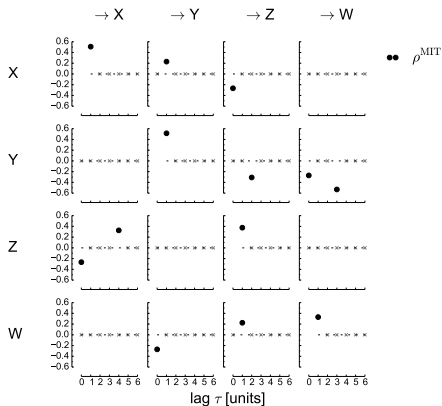
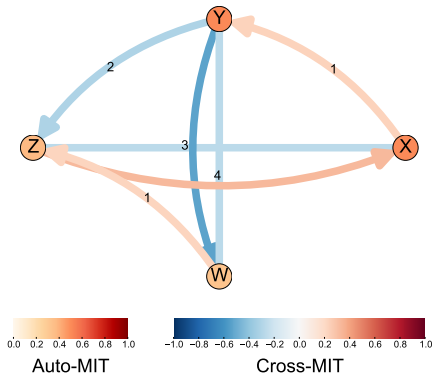


Figure : MIT plot for the simulated process. The plot of significant lags for simulated process associated with the MIT plot.

The results indicates a lagged causal chain as $X \xrightarrow{1} Y \xrightarrow{2} Z$ with feedback $Z \xrightarrow{4} X$, and $Y \xrightarrow{3} W \xrightarrow{1} Z$, plus contemporaneous links $X - Z$ and $Y - W$.

Research-in-progress

- Develop nonlinear equivalent measure of the coupling strength.
- Empirical application: Causal dependencies in financial institutions
- Further Application: Sovereign CDS, Credit Risk etc

Part II: Blind Signal Separation

Tracking dependency in large multivariate financial systems.

- **Track dependency** in large multivariate financial systems.
 - Study the **time-varying information coupling and decoupling**
 - Deduce measures of excess dependency, frailty of the market or multivariate financial risk
 - Deduce **Early Warning Indicators**.
- Build dependency **measures** for multivariate systems that exhibit :
 - Rapid changing dynamics, i.e. exhibit a high degree of non-linearity
 - Non-Gaussianity
 - Non-stationarity.

Independent Component Analysis (ICA) 😊

- Assume that the dynamics of the multivariate signal is driven by **latent independent** non-gaussian signals
- Let $\{\mathfrak{X}_t\}_{t=1}^T$ be a **multivariate process** with $d \in \mathbb{Z}^+$ components such that

$$\mathfrak{X}_t = \mathcal{A}\mathcal{S}_t \quad (11)$$

where \mathfrak{X}_t is the observed process, \mathcal{S}_t are the **unobserved** independent signals. $\widehat{\mathcal{S}}_t$ signals, as well as the **unmixing matrix** \mathcal{W} :

$$\widehat{\mathcal{S}}_t = \mathcal{W}\mathfrak{X}_t$$

- Thus $\mathcal{W} = \mathcal{A}^{-1}$ will contain all relevant **information about the dependency structure** of the system, revealed by off-diagonal elements.

- The **Idea** is to build information coupling measures entirely **based on** \mathcal{W} .
- Here we have **full information decoupling** if \mathcal{W} is the identity matrix, and **information coupling** between some components if some elements in rows are non-zero.
- Then given this measures, use rolling windows to **study how the coupling information evolves over time**. Uses it **to develop new risk measures or early warnings**.
- The mixing process \mathcal{A} **might change with time** - thus a dynamic mixing process can arise due to **regime changes** or **structural changes**.
- Note that mixing of the latent signals **does not** need to be linear.

Research continues

“If you’re not prepared to be wrong, you’ll never come up with anything original.”
(Sir Ken Robinson at TED 2006)



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Thank you for your attention 😊